

DADOS MELANOMA

Sobrevida de 205 pacientes após cirurgia de melanoma cutâneo maligno em estágio clínico I.

Resposta: Tempo até o óbito ou final do estudo, em dias/365 -> anos

Censuras: à direita tipo I (status = 1 óbito e 0 = censura)

Covariáveis: ulc = ulceração (1 = presente, 2 = ausente)

thick = espessura do tumor (1/100 mm)

sexo = sexo do paciente (0 = feminino e 1 = masculino)

Obs: log(thick) foi centrada em seu valor médio

Fonte: Drzewiecki, K.T., Ladefoged, C., and Christensen, H.E. (1980), Biopsy and prognosis for cutaneous malignant melanoma in clinical stage I. Scand. J. Plast. Reconstr. Surg. 14, 141-144.

```
library(survival)
library(timereg)
library(survivalROC)
```

```
data(melanoma)
attach(melanoma)
```

```
Status<-ifelse(status==1, 1, 0)
espessura<- log(thick) - mean(log(thick))
tempo<-days/365
esp <- ifelse(espessura > 0.05,1,0)
```

```
data1 <- as.data.frame(cbind(ulc, espessura, sex, tempo, Status, esp))
attach(data1)
```

```
set.seed(157)
n1<-dim(melanoma)[1]
ei<-rnorm(n1,0,0.001)
tempos <- data1$tempo + ei
```

```
data2 <- as.data.frame(cbind(ulc, espessura, sex, tempos, Status, esp))
detach(data1);
attach(data2)
```

Descritiva

```
> table(Status)
 0  1
148 57          # 72% CENSURAS - TIPO I
```

```
> table(ulc)
 0  1
115 90
```

```
> table(sex,Status)
      Status
sex  0  1
 0  98 28
 1  50 29
```

```
> table(ulc,Status)
      Status
ulc  0  1
 0  99 16
 1  49 41
```

```

> esp<- ifelse(espessura > 0.05,1,0)
> table(esp,Status)
  Status
esp 0 1
  0 96 13
  1 52 44

> summary(melanoma$thick)
  Min. 1st Qu.  Median    Mean 3rd Qu.    Max.
    10     97    194    292    356    1742

> summary(espessura)
  Min. 1st Qu.  Median    Mean 3rd Qu.    Max.
-2.92100 -0.64860  0.04452  0.00000  0.65160  2.23900

> summary(tempos)
  Min. 1st Qu.  Median    Mean 3rd Qu.    Max.
 0.02742  4.17500  5.49100  5.89800  8.33500 15.25000

> summary(tempos[Status==1])
  Min. 1st Qu.  Median    Mean 3rd Qu.    Max.
 0.5102  1.9670  2.9080  3.4330  4.5640  9.1440

> summary(tempos[Status==0])
  Min. 1st Qu.  Median    Mean 3rd Qu.    Max.
 0.02741  4.95900  6.10100  6.84800  8.99400 15.25000

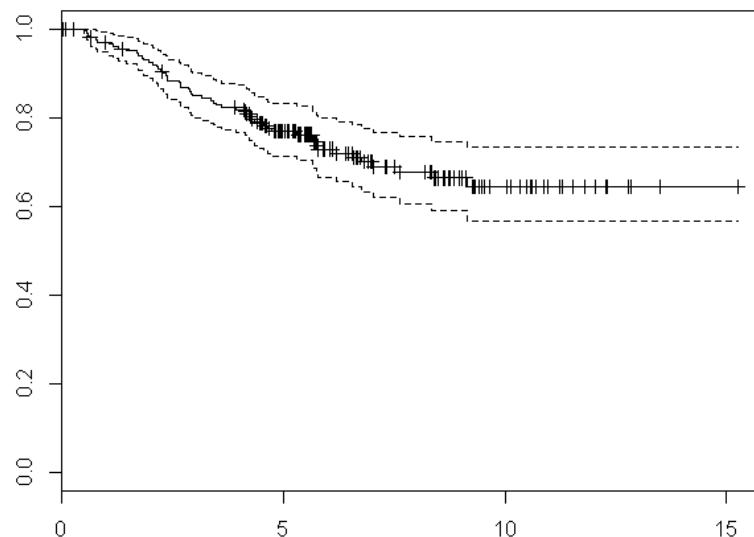
```

Kaplan-Meier Geral

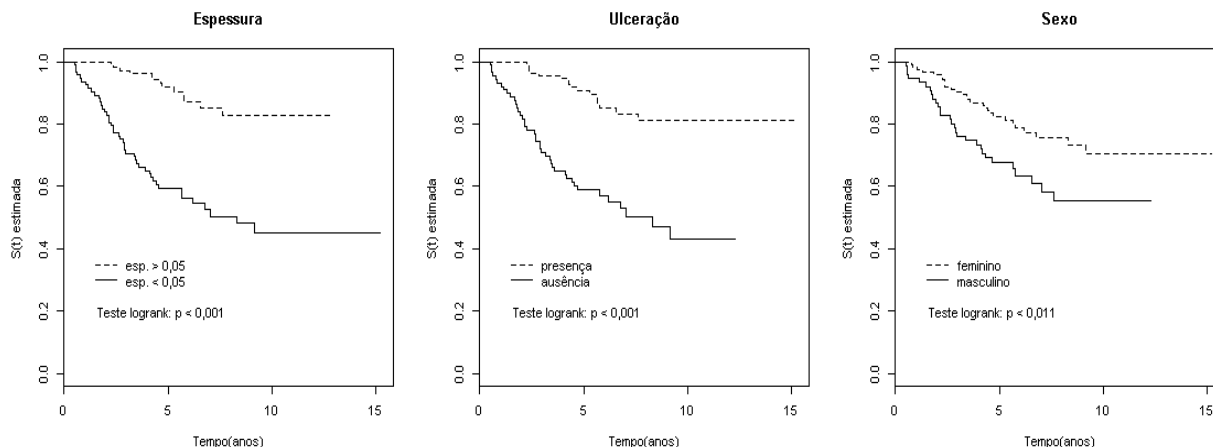
```

ekm <- survfit(Surv(tempos,Status) ~ 1, data=data2)
plot(ekm, mark.time=T)

```



Kaplan-Meier e teste logrank para as 3 covariáveis



AJUSTE MODELO DE COX

```
mod1<-coxph(Surv(tempo,Status) ~ ulc + espessura + sex, method="breslow",
data=data2)
```

	coef	exp(coef)	se(coef)	z	p
ulc	0.939	2.56	0.324	2.89	0.0038
espessura	0.576	1.78	0.179	3.21	0.0013
sex	0.381	1.46	0.271	1.41	0.1588

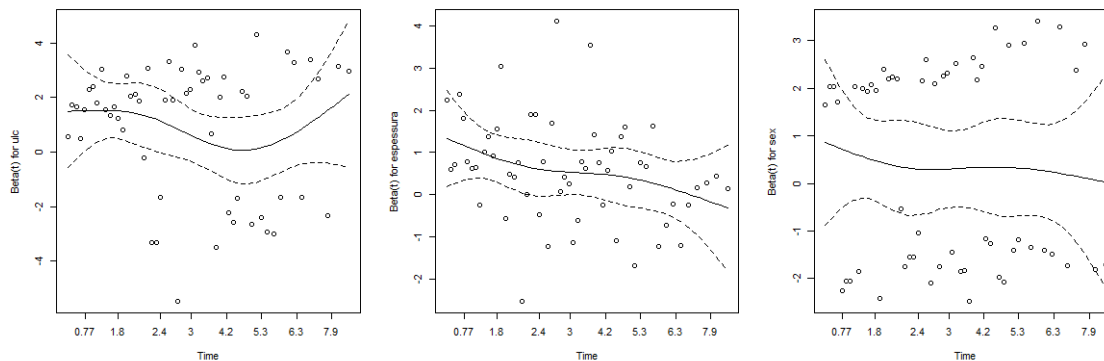
Avaliação da suposição de taxas de falha proporcionais

$$H_0: \beta_j(t) = \beta, \text{ para } j = 1, \dots, p$$

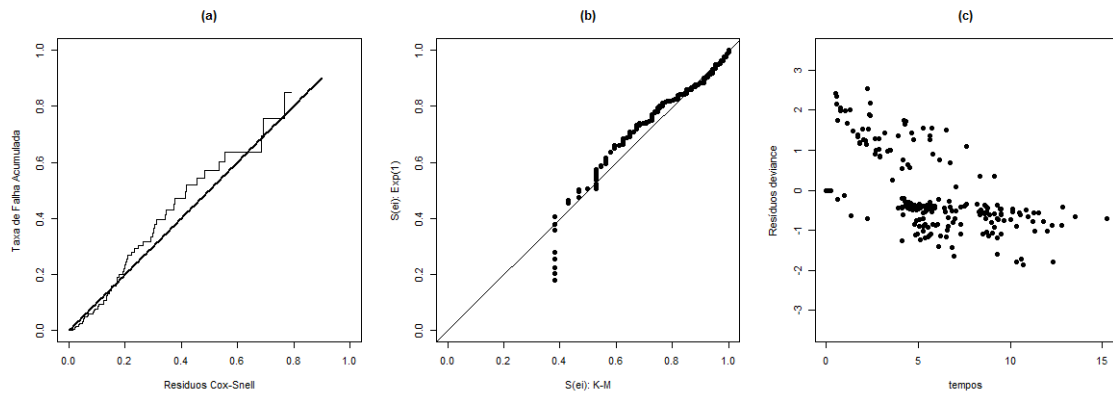
```
cox.zph(mod1)
```

	rho	chisq	p
ulc	-0.1158	0.719	0.3964
espessura	-0.2906	3.924	0.0476
sex	-0.0725	0.308	0.5792
GLOBAL	NA	8.231	0.0415

```
par(mfrow=c(1,3))
plot(cox.zph(mod1))
```



Análise dos Resíduos de Cox-Snell e Resíduos deviance



Modelo de Cox excluindo a covariável SEXO

```
mod1.1<-coxph(Surv(tempo,Status) ~ ulc + espessura, method="breslow", data=data2)
summary(mod1.1)
```

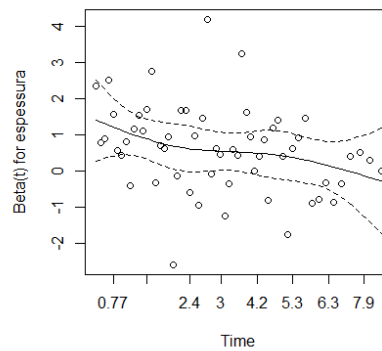
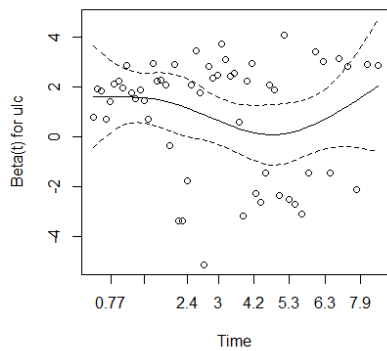
	coef	exp(coef)	se(coef)	z	Pr(> z)	
ulc	0.9712	2.6412	0.3209	3.027	0.00247	**
espessura	0.6104	1.8411	0.1759	3.470	0.00052	***

Avaliação da suposição de taxas de falha proporcionais

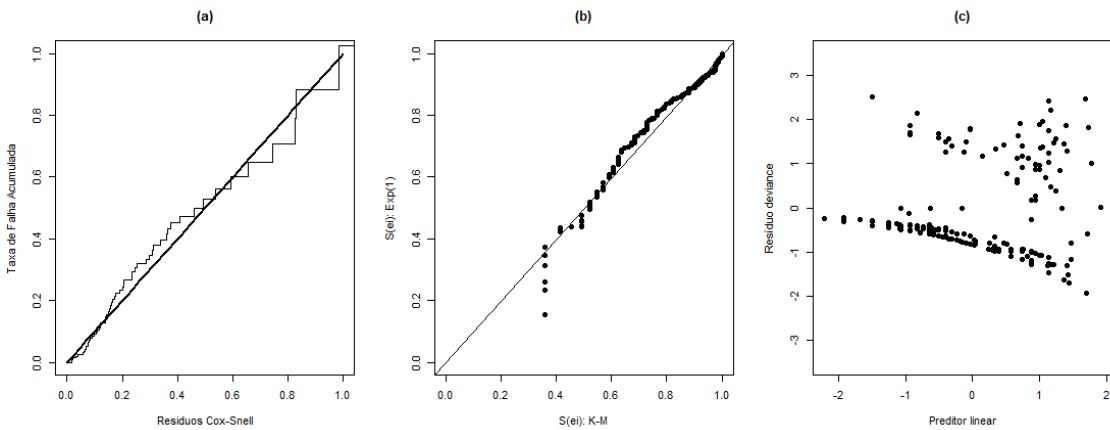
$$H_0: \beta_j(t) = \beta, \text{ para } j = 1, \dots, p$$

```
cox.zph(mod1.1)
```

	rho	chisq	p
ulc	-0.135	0.953	0.3289
espessura	-0.304	4.142	0.0418
GLOBAL	NA	7.826	0.0200



Análise dos resíduos de Cox-Snell e deviance



Modelo de Cox estratificado: espessura ~ mediana

```
esp<- ifelse(espessura > 0.05,1,0)
mod4<-coxph(Surv(tempo,Status) ~ ulc + sex + strata(esp), method="breslow",
data=data2)
```

	coef	exp(coef)	se(coef)	z	p
ulc	0.931	2.536	0.324	2.87	0.0041
sex	0.336	1.400	0.271	1.24	0.2140

Likelihood ratio test=11.4 on 2 df, p=0.00327
n= 205, number of events= 57

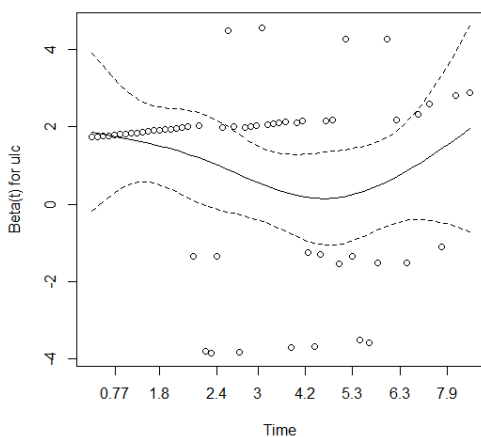
```
mod4.1<-coxph(Surv(tempo,Status) ~ ulc + strata(esp), method="breslow",
data=data2)
```

	coef	exp(coef)	se(coef)	z	p
ulc	0.961	2.615	0.321	3	0.0027

Likelihood ratio test=9.91 on 1 df, p=0.00165
n= 205, number of events= 57

```
cox.zph(mod4.1)
```

	rho	chisq	p
ulc	-0.141	1	0.316



Teste da Razão de verossimilhanças - betas comuns

```

dat1<-subset(data2,esp==0)
dat2<-subset(data2,esp==1)
mod4<-coxph(Surv(tempos,Status) ~ ulc + strata(esp), method="breslow", data=data2)
mod4.2<-coxph(Surv(tempos,Status) ~ ulc, method="breslow", data=dat1)
mod4.3<-coxph(Surv(tempos,Status) ~ ulc, method="breslow", data=dat2)
trv<- 2*(-mod4$loglik[2]+mod4.2$loglik[2]+mod4.3$loglik[2])
cbind(trv, 1-pchisq(trv,2))

```

```

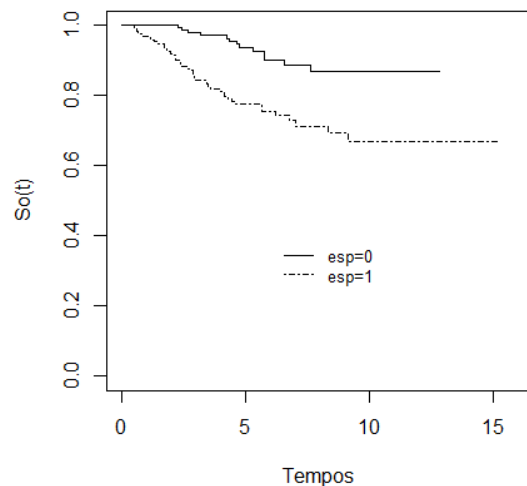
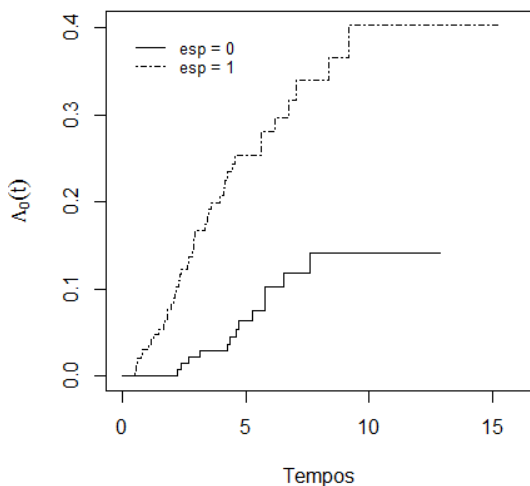
trv
[1,] 0.7278751 0.6949346

```

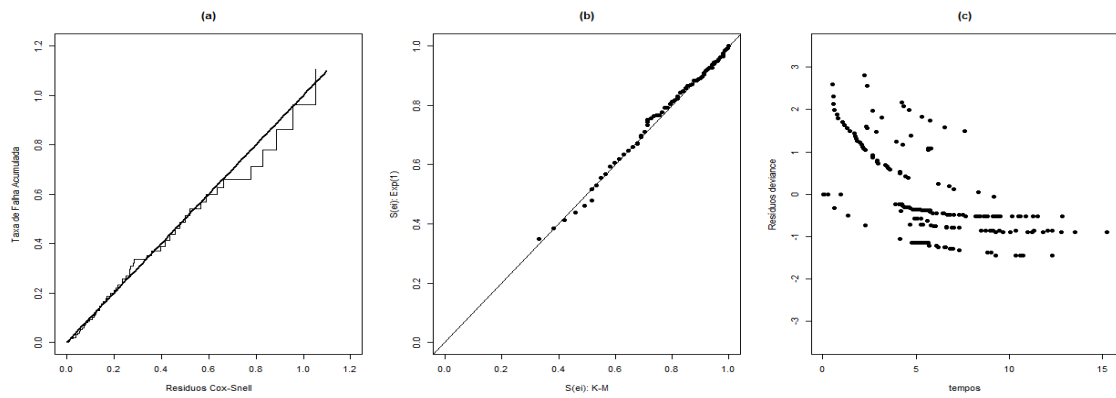
```

H0<-basehaz(mod4,centered=F) # risco acumulado de base
H01<-as.matrix(H0[1:109,1]) # estrato 1 (esp=0)
H02<-as.matrix(H0[110:205,1]) # estrato 2 (esp=1)
tempo1<- H0$time[1:109] # tempos do estrato 1
S01<-exp(-H01) # sobrevivência de base estrato 1
round(cbind(tempo1,S01,H01),digits=5) # funções de base estrato 1
tempo2<- H0$time[110:205] # tempos do estrato 2
S02<-exp(-H02) # sobrevivência de base estrato 2
round(cbind(tempo2,S02,H02),digits=5) # funções de base estrato 2
par(mfrow=c(1,2))
plot(tempo2,H02,lty=4,type="s",xlab="Tempos",xlim=range(c(0,16)),
      ylab=expression(Lambda[0]* (t)))
lines(tempo1,H01,type="s",lty=1)
legend(0.0,0.4,lty=c(1,4),c("esp = 0","esp = 1"),lwd=1,bty="n",cex=0.8)
plot(c(0,tempo1),c(1,S01),lty=1,type="s",xlab="Tempos",xlim=range(c(0,16)),
      ylim=c(0,1),ylab="So(t)")
lines(c(0,tempo2),c(1,S02),lty=4,type="s")
legend(6,0.4,lty=c(1,4),c("esp=0","esp=1"),lwd=1,bty="n",cex=0.8)

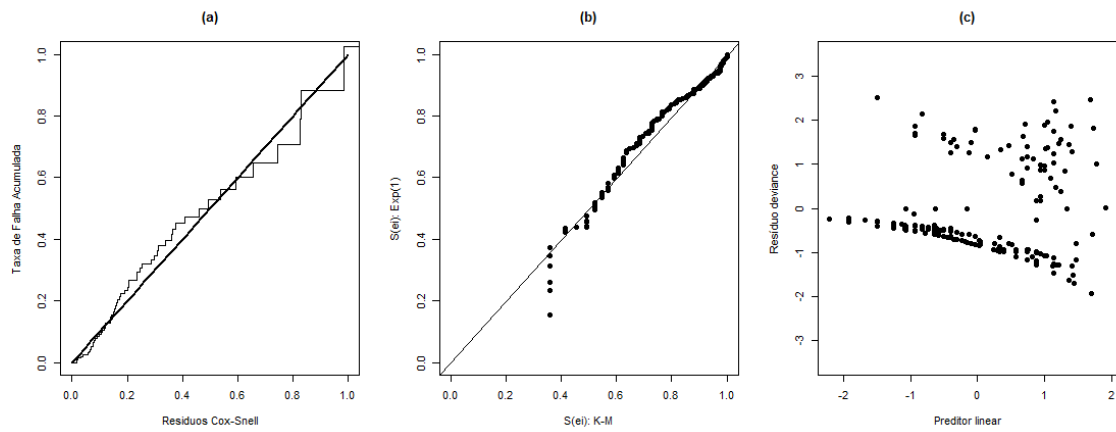
```



Análise dos Resíduos de Cox-Snell e deviance (Cox Estratificado)



Análise dos Resíduos de Cox-Snell e deviance (Cox Padrão)



Ajuste do Modelo Aditivo de Aalen

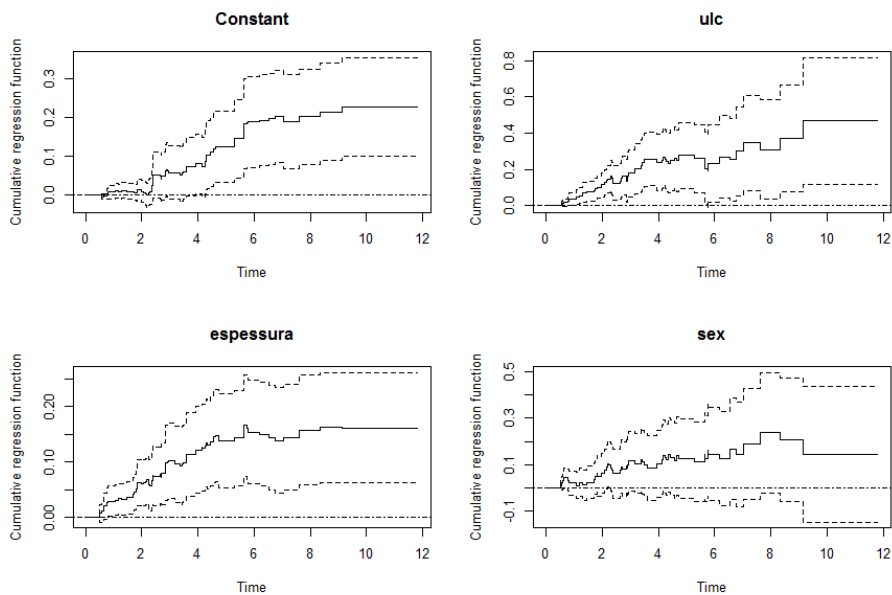
ADDREG

```
source("https://docs.ufpr.br/~giolo/Livro/ApendiceA/Addreg.r")
mod2 <- addreg(Surv(tempos,Status)~ ulc + espessura + sex, data=data2)
```

Remark: Stopped at time 12.30547 because of too low rank.
(Last estimate at time 12.27183)

Estimates at time 12.27183:

	Coef.	Std. Error	95% C.I.		Test statistic	P-vaule
Constant	0.226	0.065	0.099	0.354	3.283	0.001
ulc	0.466	0.179	0.115	0.817	3.117	0.002
espessura	0.162	0.050	0.063	0.261	3.427	0.001
sex	0.145	0.149	-0.147	0.438	1.590	0.112



TIMEREG

```
mod3 <- aalen(Surv(tempos,Status)~ ulc + espessura + sex, residuals=1,data=data2)
summary(mod3)
```

Test for non-significant effects

	Supremum-test of significance	p-value	$H_0: B(t)=0$
(Intercept)	3.33	0.014	
ulc	3.68	0.008	
espessura	3.53	0.009	
sex	2.08	0.318	

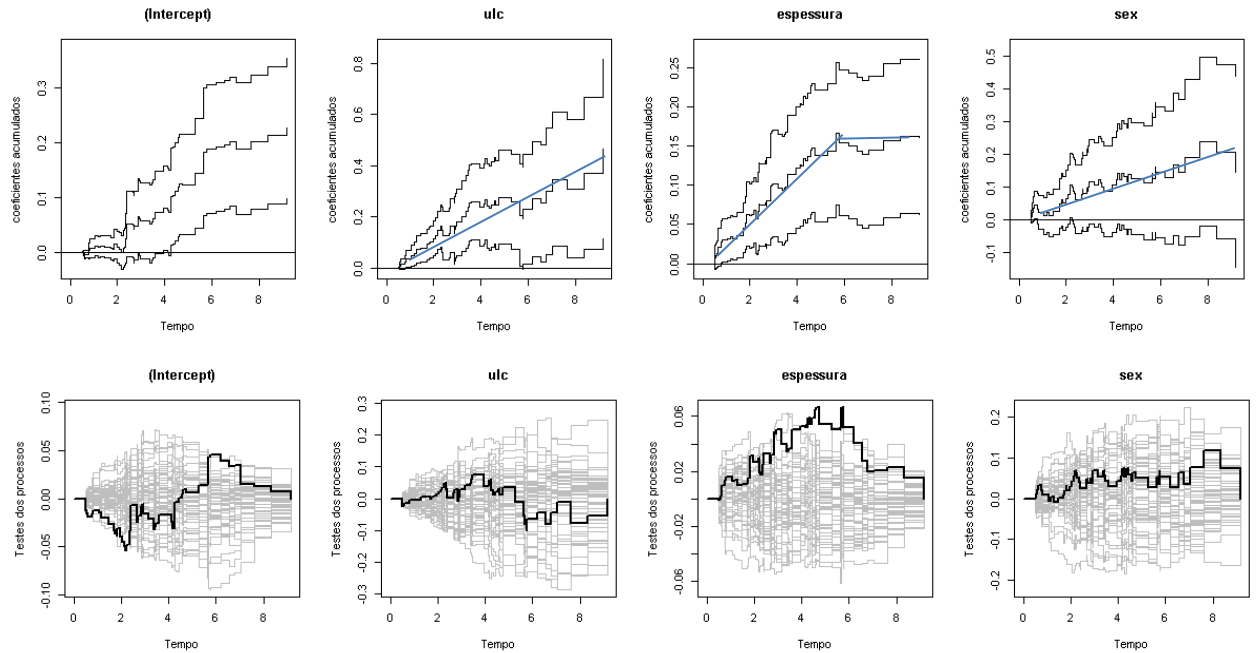
Test for time invariant effects

	Kolmogorov-Smirnov test	p-value	$H_0: \text{constant effect}$
(Intercept)	0.0538	0.332	
ulc	0.0990	0.754	
espessura	0.0662	0.049	
sex	0.1170	0.428	

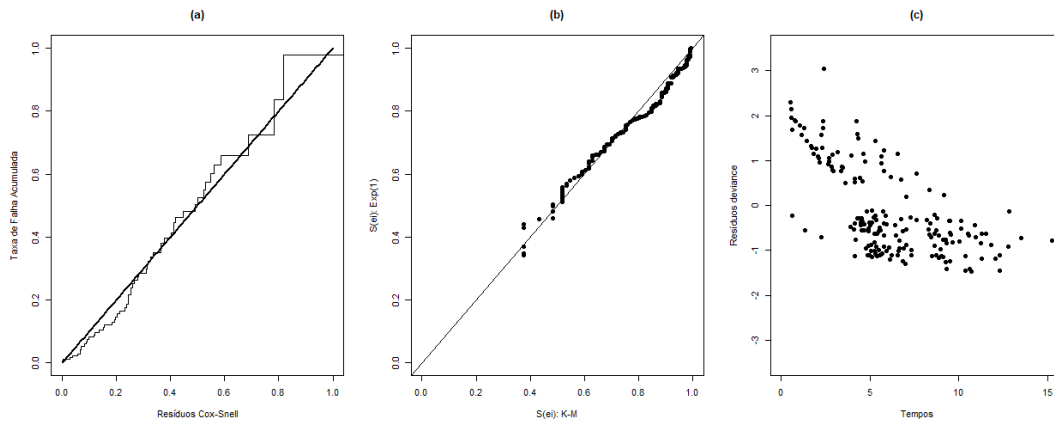
	Cramer von Mises test	p-value	$H_0: \text{constant effect}$
(Intercept)	0.00465	0.359	
ulc	0.01590	0.791	
espessura	0.01050	0.040	
sex	0.02550	0.433	

mod3\$cum

	time	(Intercept)	ulc	espessura	sex
[1,]	0.000000	0.000000e+00	0.00000000	0.00000000	0.00000000
[2,]	0.5085433	6.503850e-04	0.001322698	0.007914416	0.009926853
...					
[57,]	8.3346068	2.133328e-01	0.370026174	0.162641044	0.206156079
[58,]	9.1447507	2.263496e-01	0.466213618	0.161846932	0.145171177



Análise dos Resíduos de Cox-Snell e deviance



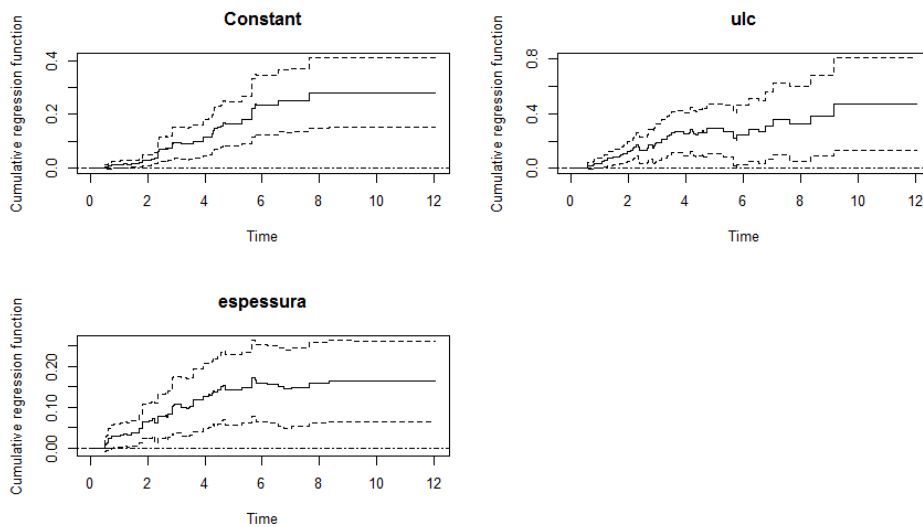
Modelo Aditivo de Aalen sem a covarável sexo

```
source("http://www.ufpr.br/~giolo/Livro/ApendiceA/Addreg.r")
mod2.1 <- addreg(Surv(tempos,Status)~ ulc + espessura, data=data2)
```

Remark: Stopped at time 12.78813 because of too low rank.
(Last estimate at time 12.30547)

Estimates at time 12.30547:

	Coef.	Std. Error	95% C.I.	Test statistic	P-value
Constant	0.281	0.066	0.151 0.411	4.579	0.000
ulc	0.470	0.173	0.130 0.809	3.224	0.001
espessura	0.163	0.050	0.064 0.261	3.474	0.001



```
mod3.1 <- aalen(Surv(tempos,Status)~ ulc + espessura, residuals=1,data=data2)
summary(mod3.1)
```

Test for non-significant effects

Supremum-test of significance p-value $H_0: B(t)=0$

(Intercept)	4.21	0.000
ulc	3.79	0.004
espessura	3.57	0.002

Test for time invariant effects

Kolmogorov-Smirnov test p-value H_0 : constant effect

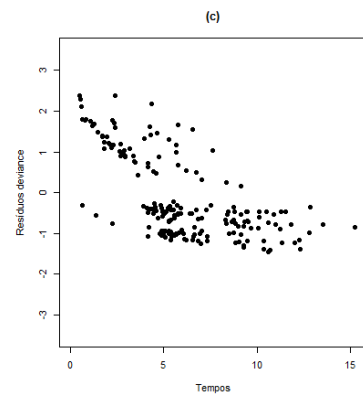
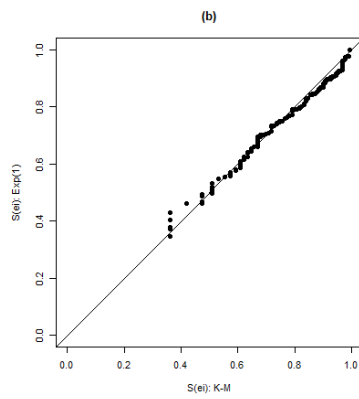
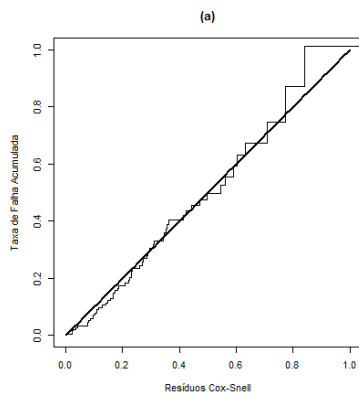
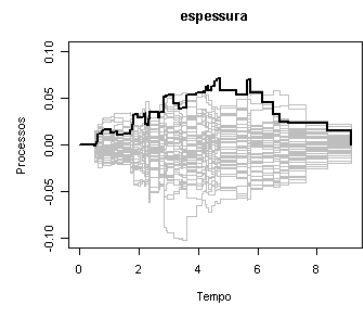
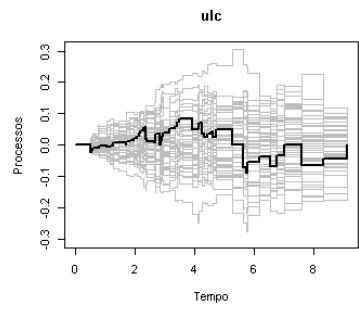
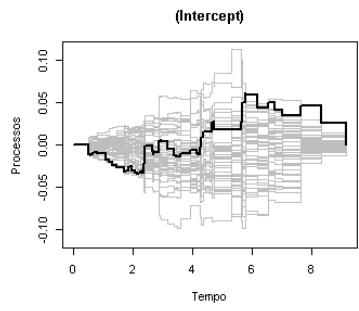
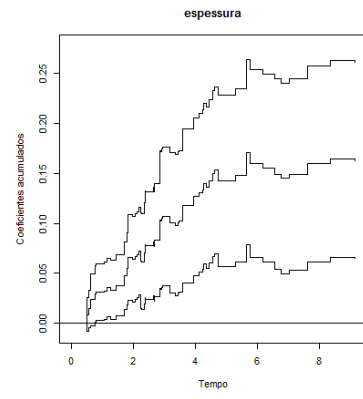
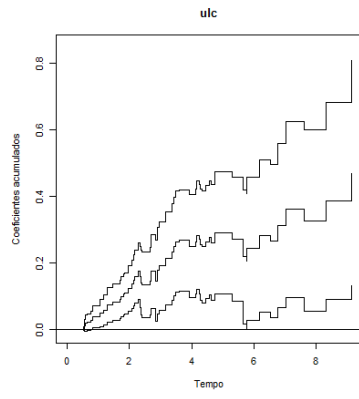
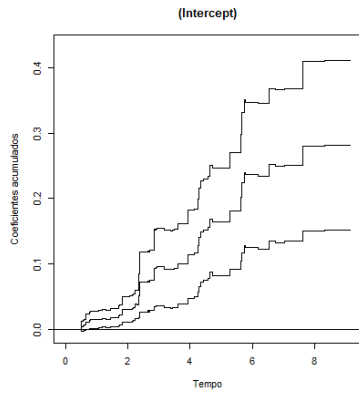
(Intercept)	0.0621	0.211
ulc	0.0907	0.800
espessura	0.0707	0.028

Cramer von Mises test p-value H_0 : constant effect

(Intercept)	0.0066	0.203
ulc	0.0144	0.786
espessura	0.0124	0.024

mod3.1\$cum

	time	(Intercept)	ulc	espessura
[1,]	0.0000000	0.000000000	0.000000000	0.000000000
[2,]	0.5085433	0.003942107	0.002367912	0.008515386
...				
[57,]	8.3346068	0.281644226	0.385255695	0.164305642
[58,]	9.1447507	0.281228405	0.469800162	0.162934266



Comparando o ajuste global dos modelos

